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NOTE

# Noise measurement from magnitude MRI using local estimates of variance and skewness

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#### Abstract

In this note, we address the estimation of the noise level in magnitude magnetic resonance (MR) images in the absence of background data. Most of the methods proposed earlier exploit the Rayleigh distributed background region in MR images to estimate the noise level. These methods, however, cannot be used for images where no background information is available. In this note, we propose two different approaches for noise level estimation in the absence of the image background. The first method is based on the local estimation of the source using maximum likelihood estimation and the second method is based on the local estimation of the skewness of the magnitude data distribution. Experimental results on synthetic and real MR image datasets show that the proposed estimators accurately estimate the noise level in a magnitude MR image, even without background data.

(Some figures in this article are in colour only in the electronic version)

## 1. Introduction

Estimation of the noise variance from magnetic resonance images (MRI) is often of key importance as an input parameter for image post-processing tasks. The estimated noise variance gives a measure of the quality of the MR data. Moreover, the noise variance is often a crucial parameter in image processing algorithms such as noise reduction, segmentation, parameter estimation or clustering (Nowak 1999, Basu *et al* 2006, Koay *et al* 2009). Many methods have been proposed in the literature for the estimation of the noise level from MRI (Brummer *et al* 1993, Nowak 1999, Sijbers and den Dekker 2004, Sijbers *et al* 2007, Aja-Fernández *et al* 2008). A survey of these methods is given in Aja-Fernández *et al* (2009). Most of the methods proposed earlier estimate the noise level from the background area of the magnitude MR image, which is known to be Rayleigh distributed (Sijbers *et al* 2007, Aja-Fernández *et al* 2008). Unfortunately, these methods cannot be used for images where no background information is available. For MR images other than the brain, like cardiac

or lung images, background data may not be available, for example in the case that the field of view (FOV) is small, such that noise assumptions based on the Rayleigh distribution fail (Aja-Fernández *et al* 2009). Also, the new scanning techniques and software eliminate most of the noisy background, which in fact affect the methods based on the Rayleigh model that need a certain amount of background pixels to perform proper estimation (Aja-Fernández *et al* 2010). Recently, in the work of Aja-Fernández *et al* (2008), a method was proposed to estimate the noise in MR images without the background region. Their method was based on the assumption that the variance of the Rician distribution approaches a Gaussian distribution at high SNR. However, at low SNR, the Rician distribution is not approximated well by a Gaussian distribution, causing a bias in the estimated noise level.

In this note, we propose two different approaches to address the aforementioned issues related to the Rician noise estimation. The first method is based on maximum likelihood (ML) estimation of the local variance for each pixel of the image using a local neighborhood. The noise variance can then be computed from the mode of distribution of the estimated local variance. The ML method simultaneously estimates the underlying true signal and noise variance from the Rician distributed data if the underlying signal is constant (Sijbers *et al* 1998). It will be shown that this method is highly accurate in estimating the noise level, but has a rather high computational complexity. Therefore, a computationally more efficient method is also proposed. This method is based on the local skewness. A correction factor for the variance estimated with the Gaussian assumption is introduced based on the estimated skewness for the actual computation of the Rician noise variance.

#### 2. Noise distribution in MRI

The raw, complex MR data acquired in the Fourier domain is characterized by a zero mean Gaussian probability density function (PDF). After the inverse Fourier transform, the noise distribution in the real and imaginary components will still be Gaussian due to the linearity and the orthogonality of the Fourier transform. However, due to the subsequent transform to a magnitude image, the noise distribution will be no longer Gaussian but Rician distributed. If *A* is the original signal amplitude, then the PDF of the reconstructed magnitude image *M* will be (Sijbers and den Dekker 2004)

$$p_{\rm M}(M|A,\sigma) = \frac{M}{\sigma^2} e^{-\frac{M^2 + A^2}{2\sigma^2}} I_0\left(\frac{AM}{\sigma^2}\right) \epsilon(M),\tag{1}$$

where  $I_0(\cdot)$  denotes the zeroth-order modified Bessel function of the first kind,  $\epsilon(\cdot)$  the Heaviside step function and  $\sigma^2$  denotes the variance of the Gaussian noise in the complex MR data. The vth raw moment of the Rice PDF can be analytically expressed as a function of the confluent hypergeometric function of the first kind, denoted by  $_1F_1$ :

$$E[M^{\nu}] = (2\sigma^2)^{\nu/2} \Gamma\left(1 + \frac{\nu}{2}\right)_1 F_1\left[-\frac{\nu}{2}; 1; -\frac{A^2}{2\sigma^2}\right],$$
(2)

where  $\Gamma$  represents the Gamma function.

When the underlying signal intensity A equals zero, the Rice PDF simplifies to a Rayleigh distribution (Henkelman 1985):

$$p_{\rm M}(M|\sigma) = \frac{M}{\sigma^2} e^{\left(-\frac{M^2}{2\sigma^2}\right)} \epsilon(M).$$
(3)

The moments of this distribution depend only on  $\sigma$  and can be written as

$$E[M^{\nu}] = (2\sigma^2)^{\nu/2} \Gamma\left(1 + \frac{\nu}{2}\right).$$
(4)

$$\widehat{\sigma_{\rm n}} = \sqrt{\frac{2}{\pi}} \bar{M}_{\rm B},\tag{5}$$

where  $\overline{M}_{\rm B}$  is the mean of the segmented background. It is also possible to estimate the noise standard deviation without an explicit segmentation by using the local statistics (Aja-Fernández *et al* 2008):

$$\widehat{\sigma_{n}} = \sqrt{\frac{2}{\pi}} \operatorname{mode}\left(\{\overline{M}_{i,j}\}\right),\tag{6}$$

where  $\overline{M}_{i,j}$  corresponds to the local mean computed for each pixel at (i, j).

At high SNR, the Rician PDF approaches a Gaussian PDF with a mean A and variance  $\sigma^2$ :

$$p_{\rm M}(M|\sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(M-A)^2}{2\sigma^2}} \epsilon(M)$$
<sup>(7)</sup>

and the moments will be

$$E[M^{\nu}] = \sigma^{2(\nu-1)} \left[ \frac{\mathrm{d}^{\nu-1}}{\mathrm{d}A^{\nu-1}} \left( A \, \mathrm{e}^{\frac{A^2}{2\sigma^2}} \right) \right] \mathrm{e}^{-\frac{A^2}{2\sigma^2}},\tag{8}$$

i.e. at high SNR, the noise variance can be estimated by simply computing the variance of a homogeneous area in the image. This approach is normally followed in the literature for the estimation of the noise in the absence of the background area. A homogeneous area in an image can be selected without manual intervention using the method suggested in Aja-Fernández *et al* (2008). The noise can be estimated from such a homogeneous area by

$$\sigma_{n}^{2} = \text{mode}\left(\left\{\sigma_{M_{i,j}}^{2}\right\}\right),\tag{9}$$

where  $\sigma_n^2$  is the variance of the estimated noise and  $\sigma_{M_i,j}^2$  corresponds to the local variance computed around each pixel at (i, j). However, at low SNR, equation (9) will lead to an underestimation of the noise level. In section 3, we discuss how to measure the noise variance from an MR image without the need for high SNR regions or a background region.

## 3. Noise estimation using local maximum likelihood estimation

Let  $m = (m_1, ..., m_N)$  be the N Rician distributed magnitude data points for which the underlying, noiseless intensity value is A. Then the joint PDF  $p_m$  can be written as (Sijbers and den Dekker 2004)

$$p_m = \prod_{n=1}^{N} \frac{M_n}{\sigma^2} e^{-\frac{M_n^2 + A^2}{2\sigma^2}} I_0\left(\frac{AM_n}{\sigma^2}\right),$$
(10)

where  $\{M_n\}$  are the magnitude variables corresponding to the magnitude observations  $\{m_n\}$ . The ML estimate of A and  $\sigma^2$  can then be constructed by substituting the available observations  $\{m_n\}$  in equation (10) and maximizing the resulting likelihood function  $L(A, \sigma^2)$  or equivalently log  $L(A, \sigma^2)$ , with respect to A and  $\sigma^2$ :

$$\ln L = \sum_{n=1}^{N} \ln\left(\frac{m_n}{\sigma^2}\right) - \sum_{n=1}^{N} \frac{m_n^2 + A^2}{2\sigma^2} + \sum_{n=1}^{N} \ln I_0\left(\frac{Am_n}{\sigma^2}\right).$$
(11)

The ML estimate is then found from the global maximum of  $\ln L$  w.r.t. A and  $\sigma^2$  (Sijbers and den Dekker 2004):

$$\left\{\widehat{A}_{\mathrm{ML}}, \widehat{\sigma^2}_{\mathrm{ML}}\right\} = \arg\{\max_{A \; \sigma^2}(\ln L)\}.$$
(12)

The above procedure assumes the underlying signal to be constant within the local neighborhood from which the signal and noise variance is estimated. If this assumption is valid for most of the local pixel neighborhoods, a robust estimator of the noise variance is given by the mode of all ML estimated local noise levels:

$$\widehat{\sigma_{\rm m}^2} = {\rm mode}(\{\widehat{\sigma_{{\rm ML}_{i,j}}^2}\}),\tag{13}$$

where  $\widehat{\sigma_m^2}$  is the estimated noise variance and  $\widehat{\sigma_{MLi,j}^2}$  is the ML estimate of the noise variance at each pixel (i, j). In fact, most of the MR images are piecewise constant with a reasonably small number of classes (e.g. MR image of brain) (Zhang *et al* 2001). Since the noise is estimated from the available piecewise constant regions in the image, this method neither depends on the image background nor on the SNR of the image.

## 4. Noise estimation using measurement of local skewness

Even though the method based on local ML estimation is accurate, a drawback is its time complexity. In this section, we propose a method based on the local computation of the skewness of the magnitude data distribution to estimate the noise variance. From equations (2), (4) and (8), it can be inferred that the moments of the Rician distribution are always in between the moments of Rayleigh and Gaussian distributions. The relationship between  $\sigma^2$  and the variance of a Rician distribution  $\sigma_M^2$  at low and high SNR can be written as

$$\sigma^2 = \sigma_{\rm M}^2 \left(2 - \frac{\pi}{2}\right)^{-1} \tag{14}$$

and

$$\sigma^2 = \sigma_{\rm M}^2,\tag{15}$$

respectively. In general,  $\sigma^2$  in terms of  $\sigma_M^2$  can be written as

Y

$$\sigma^2 = \sigma_{\rm M}^2 \times \varphi, \tag{16}$$

where  $\varphi$  is a correction factor in the range [1;  $(2 - \pi/2)^{-1}$ ], i.e. when the Rician distribution approaches a Rayleigh distribution (at low SNR), the correction factor tends to  $(2 - \pi/2)^{-1}$  and when the Rician distribution approaches a Gaussian (at high SNR), the correction factor tends to 1.

The proximity of the Rician distribution toward Rayleigh or Gaussian can be measured using its skewness, which is defined by

$$r = \frac{2E[M]^3 - 3E[M]E[M^2] + E[M^3]}{(-E[M]^2 + E[M^2])^{\frac{3}{2}}},$$
(17)

which can be analytically computed using equation (2). The skewness of a Rician distribution is a monotonically decreasing function of the SNR, defined as  $A/\sigma$  with values ranging from 0.631 (for SNR = 0) to 0 (for SNR =  $\infty$ ). If we fix A, the skewness depends only on the noise variance  $\sigma^2$  of the underlying Gaussian distribution. On the other hand, the variance of the Rician distributed data,  $\sigma_M^2$ , is given by

$$\sigma_{\rm M}^2 = E[M^2] - E[M]^2. \tag{18}$$



Figure 1. Relation between the skewness and variance correction factor for Rician distributed data.

If A is fixed,  $\sigma_M^2$  also solely depends on  $\sigma^2$ . To obtain the true value of  $\sigma^2$ ,  $\sigma_M^2$  needs to be multiplied by a correction factor:

$$\varphi = \frac{\sigma^2}{\sigma_M^2}.$$
(19)

Hence, the correction factor  $\varphi$  at various SNR can be computed by exploiting the skewness of the Rician distribution. Figure 1 shows the relationship between the correction factor  $\varphi$  and the skewness  $\gamma$  of a Rician distribution.

A lookup table can be created to represent the relationship between  $\varphi$  and  $\gamma$  for a Rician distribution from low to high SNR  $(A/\sigma)$  (i.e. from Rayleigh to Gaussian) by varying the value of A from zero to a high value and by keeping  $\sigma$  constant. The corresponding values for  $\varphi$  and  $\gamma$  can then be computed by substituting the value of A and  $\sigma$  in equations (17)–(19). Now, for every  $\gamma$ , a corresponding value for  $\varphi$  can be found from the lookup table. A polynomial model that computes an approximation of the lookup table is given in the appendix.

Now, to compute the noise variance in an MR image, the sample skewness and sample variance are computed for each pixel (i, j) within a neighborhood  $\Omega$ . The correction factor for each pixel can then be found from the lookup table and the noise variance for each pixel is estimated using equation (16). Following a similar reasoning as in the previous section, the noise variance can now be estimated as the mode of all local estimates of noise variance around each pixel, which can be written as

$$\widehat{\sigma}_{\mathsf{S}}^2 = \mathsf{mode}(\{\widehat{\sigma}_{\mathsf{L}_{i,j}}^2\}),\tag{20}$$

where  $\widehat{\sigma}_{s}^{2}$  is the estimate of the noise variance in the image and  $\widehat{\sigma}_{L_{i,j}}^{2}$  is the estimated noise variance for a neighborhood around pixel (i, j).

#### 5. Experimental results

Experiments were conducted on both real and synthetic MR images with and without the background region. The window size used in our experiments for local estimates was  $9 \times 9$ . For simulations we used the standard MR image phantom of the brain obtained from the Brainweb database (Cocosco *et al* 1997) and a cardiac MR image both with intensity values in the range 0–255. The cardiac image, in contrast to the brain image, contained no background areas. Both images were artificially corrupted with Rician noise with the noise level  $\sigma$  as 50. Figure 2 shows the distribution of the local estimates of  $\sigma$  computed using ML and skewness-based methods for both the brain and the cardiac images. It can be observed from the distributions that their mode is almost equal to the true standard deviation of the noise.



**Figure 2.** Noisy MR images of the brain (a) and heart (c), both corrupted with Rician noise with  $\sigma = 50$ . On the right, the distribution of local estimates of  $\sigma$  computed using both ML (solid red line) and skewness-based methods (dashed blue line) are shown for both the brain (b) and the cardiac (d) images.

This experiment demonstrates the independence of the proposed estimators to the image background.

Now, to show the reliability of the proposed methods on both low and high SNR, we conducted the experiments on the cardiac MR image after corrupting the image with Rician noise with  $\sigma$  ranging from 10 to 100. Results of this experiment are shown in figure 3. The mean of 100 experiments divided by the actual value of  $\sigma$  is depicted. The value closer to 1 is the best estimate. The results are compared with the estimator given in equation (9). Since, at high SNR the Rician PDF approaches a Gaussian PDF, the estimation based on equation (9) will be closer to the true  $\sigma$ . However, as SNR drops, as expected, it can be seen from the figure that the Gaussian assumption introduces a bias. It can also be observed from the graph that the proposed estimators are significantly less biased especially for low values of the SNR, the proposed estimator based on the local ML estimation of  $\sigma$  is closer to the true  $\sigma$  than the one based on the skewness. However, the time complexity of the skewness-based method is less than that of the ML-based method.

For the experiments on real data we used the MR image of a cherry tomato (with a mean intensity of around 10 000). A set of MR images was reconstructed by averaging 1 to 12 images. Averaging was done in the complex k-space. The resulting noise variance



**Figure 3.** The estimated  $\sigma$ /actual  $\sigma$  for various values of  $\sigma$  ranging from 10 to 100. This experiment was done on the cardiac MR image. The red and blue lines correspond to the proposed methods using ML and skewness, respectively, and the green line corresponds to the estimator given in equation (9). The shaded area shows the standard deviation for 100 experiments.



**Figure 4.** Estimated  $\sigma$  as a function of the number of averages *n* used during the acquisition. MR image of a cherry tomato was used for this experiment.

as a function of the number of averages over *n* images was then estimated. The theoretical reduction of the noise standard deviation as a function of the number of images *n* over which the average was taken is known to be  $1/\sqrt{n}$ . Since the experimental setup for all acquisitions was identical except for averaging, the estimated noise standard deviation  $\hat{\sigma}$  multiplied by  $\sqrt{n}$  is expected to be constant as a function of *n*. It can be seen from figure 4 that the proposed method exhibits this property.

# 6. Conclusion

Two different approaches based on the local estimates of the variance and skewness were proposed to address the estimation of the noise from MR images when the Rayleigh background data or high SNR image regions are not available. The estimation of  $\sigma$  based on local ML is slightly more accurate than the method based on the skewness. However, the time complexity of the ML-based method is significantly greater than that of the skewness-based method. Experimental results on synthetic and real MR images show the reliability of the proposed methods. Even though the proposed methods are based on the noise characteristics of magnitude image data acquired with single coil MRI, they can, under certain conditions, be extended for multiple coil imaging methods like SENSE.

Table A1. Val	ues of the coeffi	cients in equation (A.1).
Coefficient	Value	-
$p_0$	1	
$p_1$	2.89	
$p_2$	$-7.29 \times 10$	
$p_3$	$1.16 \times 10^{3}$	
$p_4$	$-9.84 \times 10^{3}$	
$p_5$	$4.78 \times 10^{4}$	
$p_6$	$-1.38 \times 10^{5}$	
$p_7$	$2.31 \times 10^{5}$	
$p_8$	$-2.09 \times 10^{5}$	
$p_9$	$7.86 \times 10^{4}$	

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# Appendix. Expression for the lookup table

The lookup table mentioned in section 4 can be approximated by the following polynomial equation:

$$\varphi(\gamma) = \sum_{i=0}^{9} p_i \gamma^i, \tag{A.1}$$

where  $\varphi$  is the correction factor and  $\gamma$  is the skewness. The values for the coefficients (with 95% confidence interval) are given in table A1.

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